

Recovery Failure Probability of Power-based NOMA on the Uplink of a 5G Cell for an Arbitrary Number of Superimposed Signals

Maria Luisa Merani

Dipartimento di Ingegneria “Enzo Ferrari”

University of Modena and Reggio Emilia, Modena, Italy

e-mail:marialuisa.merani@unimore.it

Abstract—This work puts forth an analytical approach to evaluate the recovery failure probability of power-based Non-Orthogonal Multiple Access (NOMA) on the uplink of a 5G cell, the recovery failure being defined as the unfortunate event where the receiver is unable to decode even one out of the n simultaneously received signals. In the examined scenario, Successive Interference Cancellation (SIC) is considered and an arbitrary number of superimposed signals is present. For the Rayleigh fading case, the recovery failure probability is provided in closed-form, clearly outlining its dependency on the signal-to-noise ratio of the users that are simultaneously transmitting, as well as on their distance from the receiver.

keywords - Uplink NOMA, successive interference cancellation, recovery failure probability, 5G

I. INTRODUCTION

In recent years NOMA has stirred much interest, because of its promise of enhancing the capacity of 5G cellular systems. NOMA allows different simultaneous users to share the available system resources (frequency, time) through a variety of different techniques, well illustrated in [1] and [2]: as a matter of fact, NOMA can operate in the power-domain, can adopt spreading sequences, can rely on coding matrices and/or interleaving. This work focuses on power-based NOMA on the uplink of a 5G cell, when a SIC receiver is employed. In [3], emphasis was on uplink power-based NOMA too, the goal being to evaluate the achievable sum data rate and the corresponding outage, which was provided in closed-form for the case of two users. The latter condition is commonly encountered in literature, as

it guarantees a tractable analysis. Unlike [3] and previous works, this study sets no limit to the number of superimposed signals. Its main original contribution consists in the introduction of an analytical framework to evaluate the recovery failure probability, which is defined as the probability that the recovery of the signals simultaneously arriving at the receiver fails. In other words, not even one out of the n received signals can be successfully decoded; this condition happens when the constraints on the received powers that the SIC receiver mandates are not observed. For the case of Rayleigh fading, the theoretical analysis provides the recovery failure probability in closed-form, immediately revealing its dependency on several factors, among which frequency, different signal-to-noise ratio assignments and distance of the users from the receiver. The numerical results explore the setting where either two or three superimposed signals are present, and show that several configurations exist, where the recovery failure probability is confined to low values and the benefit of NOMA can indeed be effectively exploited.

The remainder of this manuscript is organized as follows: Section II defines the scenario of investigation and develops the theoretical analysis; the case of Rayleigh fading is illustrated in Section III; Section IV provides some reference results, while the conclusions are drawn in Section V.

II. SCENARIO AND PERFORMANCE ANALYSIS

Within the current work, uplink communications in a 5G cell are examined; power-based NOMA is employed and the reference scenario features n User

Equipments (*UEs*) that transmit to the enhanced node B (enodeB), on the same radio spectrum. Let h_i denote the envelope of the channel between the i -th *UE*, UE_i , $i = 1, 2, \dots, n$, and the enodeB; as in [4], let

$$\Gamma_i = \frac{h_i^2}{N_0 B} \quad (1)$$

indicate the i -th normalized channel gain, N_0 being the noise power spectral density and B the transmission bandwidth. Further, let $p_{\Gamma_i}(\gamma_i)$ be the probability density function (pdf) of the generic Γ_i , $0 \leq \gamma_i \leq +\infty$. The assumption is that, $\forall i$ and j , Γ_i and Γ_j be independent random variables with different mean values, that is to say, different *UEs* experience independent channel conditions during their transmissions to the enodeB.

Unlike LTE, 5G uplink power-based NOMA mandates for different transmit power levels for different *UEs*, in order to guarantee distinct received powers at the enodeB and to facilitate the task of the interference cancellation receiver. Moreover, unlike downlink power-based NOMA [5], which assigns users with stronger channel conditions lower power levels, for uplink NOMA the more favorable the channel gain that the *UE* experiences, the higher the transmit power level that the *UE* has to work with. In other words, before NOMA *UEs* transmit, the enodeB ranks them according to their channel gains, strongest first; then, the *UE* with the largest channel gain is assigned the highest transmit power, the *UE* with the second largest gain the second highest power, until the least favored *UE*, which operates with the lowest transmit power level.

No matter what power assignment law is adopted, the instantaneous values of the Γ_i s need to be ordered, so as to obtain $\Gamma_{(1)}, \Gamma_{(2)}, \dots, \Gamma_{(n)}$, with

$$\Gamma_{(1)} > \Gamma_{(2)} > \dots > \Gamma_{(n)}; \quad (2)$$

next, transmit powers are assigned to *UE*'s respecting the constraint

$$P_{(1)} > P_{(2)} > \dots > P_{(n)}, \quad (3)$$

$P_{(i)}$ representing the transmit power of $UE_{(i)}$, the *UE* with the i -th largest channel gain $\Gamma_{(i)}$.

Note that in the ordered set $\Gamma_{(1)}, \Gamma_{(2)}, \dots, \Gamma_{(n)}$ the new random variables $\Gamma_{(i)}$ s are no longer independent and that, for the notation employed, $\Gamma_{(1)}$ is the largest order statistic, whereas $\Gamma_{(n)}$ is the smallest.

The enodeB receives the superimposed messages from the *UEs* and through successive interference cancellation it attempts to decode them: if totally successful, it recovers all n signals, the strongest first, then the second strongest, until the weakest.

We begin observing that the strongest signal received by enodeB from $UE_{(1)}$ can be detected if the difference between the power received from $UE_{(1)}$ and the sum of the powers received from all other users is greater than or equal to a threshold value P_{thres} , that represents the minimum difference required to extract the first signal. In analytical terms, the condition of successful recovery for such signal is given by

$$P_{(1)}\Gamma_{(1)} - \sum_{i=2}^n P_{(i)}\Gamma_{(i)} \geq P_{thres}. \quad (4)$$

If (4) is satisfied, then the first decoding event occurs, the SIC receiver recovers one signal and next proceeds and attempts to decode the second strongest signal, received from $UE_{(2)}$. The first signal is subtracted out, so that the recovery of the second strongest received signal requires the condition

$$P_{(2)}\Gamma_{(2)} - \sum_{i=3}^n P_{(i)}\Gamma_{(i)} \geq P_{thres} \quad (5)$$

to hold. If (5) is verified, then the receiver decodes two signals and can tackle the decoding of the third strongest signal; on the contrary, if last condition is not fulfilled, the SIC receiver halts. In general, the process stops when either the last, weakest signal is successfully decoded, or when k signals only, $k < n - 1$, are recovered, this event corresponding to the occurrence of the condition:

$$P_{(k+1)}\Gamma_{(k+1)} - \sum_{i=k+2}^n P_{(i)}\Gamma_{(i)} < P_{thres} \quad (6)$$

The final consequence of this iterated reasoning is that inequality (4) represents the necessary and sufficient condition for the decoding of the strongest signal, but also the necessary condition for the decoding of 2, 3, $\dots$, n signals out of n .

Next, let us define P_{fail_n} , the recovery failure probability of power-based uplink NOMA in the presence of n simultaneous transmissions, as the probability that not even one, out of the n superimposed signals, can be correctly recovered by the

SIC receiver. From the previous reasoning it follows that

$$\begin{aligned} P_{fail_n} &= \\ &= 1 - Pr\{\text{the strongest signal is recovered}\} = \\ &= Pr\{P_{(1)}\Gamma_{(1)} - \sum_{i=2}^n P_{(i)}\Gamma_{(i)} < P_{thres}\}. \end{aligned} \quad (7)$$

Evaluating (7) requires the consideration of n dependent random variables, the generic of which is

$$X_{(i)} = P_{(i)}\Gamma_{(i)}; \quad (8)$$

recalling (2) and (3), it is immediate to conclude that condition $X_{(1)} > X_{(2)} > \dots > X_{(n)}$ holds.

To understand how the P_{fail_n} evaluation can be pursued, it is instructive to consider $n = 2$ and to indicate by $f_{joint_2}(x_{(1)}, x_{(2)})$ the joint pdf of the random variables $X_{(1)}$ and $X_{(2)}$, so that the recovery failure probability in (7) becomes

$$\begin{aligned} P_{fail_2} &= Pr\{P_{(1)}\Gamma_{(1)} - P_{(2)}\Gamma_{(2)} < P_{thres}\} = \\ &= Pr\{X_{(1)} - X_{(2)} < P_{thres}\} = \\ &= 1 - Pr\{X_{(2)} < X_{(1)} - P_{thres}\} = \\ &= 1 - \int_{x_{(1)}=P_{thres}}^{\infty} \int_{x_{(2)}=0}^{x_{(1)}-P_{thres}} f_{joint_2}(x_{(1)}, x_{(2)}) \cdot \\ &\quad \cdot dx_{(2)} dx_{(1)}. \end{aligned} \quad (9)$$

For a generic n , last expression generalizes to

$$\begin{aligned} P_{fail_n} &= 1 - \\ &\int_{x_{(1)}=P_{thres}}^{\infty} \int_{x_{(2)}=0}^{x_{(1)}-P_{thres}} \int_{x_{(3)}=0}^{x_{(1)}-x_{(2)}-P_{thres}} \dots \\ &\int_{x_{(n)}=0}^{x_{(1)}-x_{(2)}-x_{(3)}-\dots-x_{(n-1)}-P_{thres}} \\ &\quad f_{joint_n}(x_{(1)}, x_{(2)}, \dots, x_{(n)}) \cdot \\ &\quad \cdot dx_{(n)} dx_{(n-1)} \dots dx_{(1)}, \end{aligned} \quad (10)$$

where $f_{joint_n}(x_{(1)}, x_{(2)}, \dots, x_{(n)})$ indicates the joint pdf of the ordered set $X_{(1)}, X_{(2)}, \dots, X_{(n)}$.

The next problem at hand is to determine such pdf. In this respect, let $f_i(\cdot)$ be the pdf of the – *unordered* – random variable X_i , defined as

$$X_i = P_{(i)} \cdot \Gamma_i, \quad i = 1, 2, \dots, n, \quad (11)$$

whose pdf is immediately determined, once $p_{\Gamma_i}(\gamma_i)$ is known, as $P_{(i)}$ is a constant, and define F_n as the following $n \times n$ matrix

$$F_n = \begin{bmatrix} f_1(x_{(1)}) & f_2(x_{(1)}) & \dots & f_n(x_{(1)}) \\ f_1(x_{(2)}) & f_2(x_{(2)}) & \dots & f_n(x_{(2)}) \\ \vdots & \vdots & \ddots & \vdots \\ f_1(x_{(n)}) & f_2(x_{(n)}) & \dots & f_n(x_{(n)}) \end{bmatrix}. \quad (12)$$

For the purpose of what follows, recall that the permanent of a square matrix A , written as $|A|$, is defined like the determinant, except that all signs are positive. Again, to make things clearer, consider the case $n = 2$ and observe that in this case the permanent of F_2 is given by:

$$\begin{aligned} |F_2| &= \begin{vmatrix} f_1(x_{(1)}) & f_2(x_{(1)}) \\ f_1(x_{(2)}) & f_2(x_{(2)}) \end{vmatrix} = \\ &= f_1(x_{(1)})f_2(x_{(2)}) + f_2(x_{(1)})f_1(x_{(2)}). \end{aligned} \quad (13)$$

For an arbitrary n , it can be demonstrated that the joint pdf $f_{joint_n}(x_{(1)}, x_{(2)}, \dots, x_{(n)})$ of the ordered statistics $X_{(1)}, X_{(2)}, \dots, X_{(n)}$ is

$$f_{joint_n}(x_{(1)}, x_{(2)}, \dots, x_{(n)}) = |F_n|, \quad (14)$$

F_n being given by (12). Last result is substantiated by the reasoning in [6] and [7], the arguments of [8] being extended to prove the formulation in (14) with the use of permanents.

At first sight, $f_{joint_n}(x_{(1)}, x_{(2)}, \dots, x_{(n)})$ gives the impression that evaluating the integral in (10) might be quite cumbersome when n is arbitrary. However, the joint pdf obeys a highly peculiar structure and an alike – and more convenient – rewriting of it is provided in the following terms: let $\mathcal{S}_N$ indicate all $n!$ permutations of the set $N = \{1, 2, \dots, n\}$ and by $S_i = \{i_1, i_2, \dots, i_n\}$ the generic of such permutations. It follows that $f_{joint_n}(x_{(1)}, x_{(2)}, \dots, x_{(n)})$ is equivalently written as

$$\begin{aligned} f_{joint_n}(x_{(1)}, x_{(2)}, \dots, x_{(n)}) &= \\ &= \sum_{S_i \in \mathcal{S}_N} f_1(x_{(i_1)}) f_2(x_{(i_2)}) \cdot \dots \cdot f_n(x_{(i_n)}). \end{aligned} \quad (15)$$

Last expression highlights that the joint pdf exhibits the presence of $n!$ terms, wherein the permutations of the arguments of the $f_1(\cdot), f_2(\cdot), \dots, f_n(\cdot)$ pdfs appear. Replacing (15) in (10) gives P_{fail_n} as

$$\begin{aligned} P_{fail_n} &= 1 - \\ &\int_{x_{(1)}=P_{thres}}^{\infty} \int_{x_{(2)}=0}^{x_{(1)}-P_{thres}} \int_{x_{(3)}=0}^{x_{(1)}-x_{(2)}-P_{thres}} \dots \\ &\int_{x_{(n)}=0}^{x_{(1)}-x_{(2)}-x_{(3)}-\dots-x_{(n-1)}-P_{thres}} \\ &\quad \sum_{S_i \in \mathcal{S}_N} f_1(x_{(i_1)}) f_2(x_{(i_2)}) \cdot \dots \cdot f_n(x_{(i_n)}) \cdot \\ &\quad \cdot dx_{(n)} dx_{(n-1)} \dots dx_{(1)}, \end{aligned} \quad (16)$$

and indicating by I_{S_i} the result of the integral

$$I_{S_i} = \int_{x_{(1)}=P_{thres}}^{\infty} \int_{x_{(2)}=0}^{x_{(1)}-P_{thres}} \int_{x_{(3)}=0}^{x_{(1)}-x_{(2)}-P_{thres}} \dots \int_{x_{(n)}=0}^{x_{(1)}-x_{(2)}-x_{(3)}-\dots-x_{(n-1)}-P_{thres}} p_{i_1 i_2 \dots i_n}(x_{(1)}, x_{(2)}, \dots, x_{(n)}) \cdot dx_{(n)} dx_{(n-1)} \dots dx_{(1)}, \quad (17)$$

where

$$p_{i_1 i_2 \dots i_n}(x_{(1)}, x_{(2)}, \dots, x_{(n)}) = f_1(x_{(i_1)}) f_2(x_{(i_2)}) \dots f_n(x_{(i_n)}), \quad (18)$$

then P_{fail_n} is rewritten as

$$P_{fail_n} = 1 - \sum_{S_i \in \mathcal{S}_N} I_{S_i}. \quad (19)$$

Luckily, when the random variables $X_1, X_2, \dots, X_n$ obey the same statistical description, although with different mean values, for a permutation S_j different than S_i , the I_{S_j} result is readily obtained from I_{S_i} through the analogous permutation of the $f_i(\cdot)$'s arguments $x_{(i_k)}$, $k = 1, 2, \dots, n$, in (18), $\forall j$. That is to say, given the n -th fold integral in (17) has been solved once, e.g., I_{S_1} has been determined, $S_1 = \{1, 2, \dots, n\}$, then all the remaining I_{S_i} terms are known. This significantly reduces P_{fail_n} computational complexity in n , no matter what statistical description is examined for Γ_i , $i = 1, 2, \dots, n$.

As an illustrative example, the case of Rayleigh fading is examined in next Section.

III. RAYLEIGH FADING CASE

When the envelope of the received signal is subject to Rayleigh fading, Γ_i and in turn X_i are exponentially distributed with means $\bar{\Gamma}_i$ and $\bar{X}_i$, respectively. Beginning with the case $n = 2$, from (19) P_{fail_2} specializes to

$$P_{fail_2} = 1 - (I_{S_1} + I_{S_2}), \quad (20)$$

where $S_1 = \{1, 2\}$ and

$$I_{S_1} = \int_{x_{(1)}=P_{thres}}^{\infty} \int_{x_{(2)}=0}^{x_{(1)}-P_{thres}} p_{12}(x_{(1)}, x_{(2)}) dx_{(2)} dx_{(1)} \quad (21)$$

with

$$p_{12}(x_{(1)}, x_{(2)}) = f_1(x_{(1)}) f_2(x_{(2)}) = \frac{1}{\bar{X}_1} \exp\left(-\frac{x_{(1)}}{\bar{X}_1}\right) \cdot \frac{1}{\bar{X}_2} \exp\left(-\frac{x_{(2)}}{\bar{X}_2}\right); \quad (22)$$

analogously, $S_2 = \{2, 1\}$ and

$$I_{S_2} = \int_{x_{(1)}=P_{thres}}^{\infty} \int_{x_{(2)}=0}^{x_{(1)}-P_{thres}} p_{21}(x_{(1)}, x_{(2)}) dx_{(2)} dx_{(1)} \quad (23)$$

with

$$p_{21}(x_{(1)}, x_{(2)}) = f_1(x_{(2)}) f_2(x_{(1)}) = \frac{1}{\bar{X}_1} \exp\left(-\frac{x_{(2)}}{\bar{X}_1}\right) \cdot \frac{1}{\bar{X}_2} \exp\left(-\frac{x_{(1)}}{\bar{X}_2}\right). \quad (24)$$

Solving the integral in (21) gives

$$I_{S_1} = \frac{\bar{X}_1}{\bar{X}_1 + \bar{X}_2} \exp\left(-\frac{P_{thres}}{\bar{X}_1}\right), \quad (25)$$

wherefore I_{S_2} immediately follows as

$$I_{S_2} = \frac{\bar{X}_2}{\bar{X}_2 + \bar{X}_1} \exp\left(-\frac{P_{thres}}{\bar{X}_2}\right), \quad (26)$$

and finally

$$P_{fail_2} = 1 + \left(\frac{\bar{X}_1}{\bar{X}_1 + \bar{X}_2} \exp\left(-\frac{P_{thres}}{\bar{X}_1}\right) + \frac{\bar{X}_2}{\bar{X}_1 + \bar{X}_2} \exp\left(-\frac{P_{thres}}{\bar{X}_2}\right) \right). \quad (27)$$

Similarly, when $n = 3$, there will be 3! distinct integral contributions of the type in (17) in the recovery failure probability expression, that are determined once $p_{123}(x_{(1)}, x_{(2)}, x_{(3)})$ is introduced,

$$p_{123}(x_{(1)}, x_{(2)}, x_{(3)}) = \frac{1}{\bar{X}_1} \exp\left(-\frac{x_{(1)}}{\bar{X}_1}\right) \cdot \frac{1}{\bar{X}_2} \exp\left(-\frac{x_{(2)}}{\bar{X}_2}\right) \cdot \frac{1}{\bar{X}_3} \exp\left(-\frac{x_{(3)}}{\bar{X}_3}\right). \quad (28)$$

Now, $S_1 = \{1, 2, 3\}$ and

$$I_{S_1} = \int_{x_{(1)}=P_{thres}}^{\infty} \int_{x_{(2)}=0}^{x_{(1)}-P_{thres}} \int_{x_{(3)}=0}^{x_{(1)}-x_{(2)}-P_{thres}} p_{123}(x_{(1)}, x_{(2)}, x_{(3)}) \cdot dx_{(3)} dx_{(2)} dx_{(1)}, \quad (29)$$

that after a few passages is determined as

$$I_{S_1} = \frac{\bar{X}_1^2 \exp(-\frac{P_{thres}}{\bar{X}_1})}{(\bar{X}_1 + \bar{X}_2)(\bar{X}_1 + \bar{X}_3)}. \quad (30)$$

Moreover, it can be verified that the set $S_2 = \{1, 3, 2\}$ gives the same result as S_1 , that is to say, $I_{S_1} = I_{S_2}$. If we now introduce $S_3 = \{2, 1, 3\}$ and $S_4 = \{2, 3, 1\}$, suitably permuting $\bar{X}_1, \bar{X}_2$ and $\bar{X}_3$ in (30) $I_{S_3} = I_{S_4}$ is also determined; the same applies

to $S_5 = \{3, 2, 1\}$ and $S_6 = \{3, 1, 2\}$, for which $I_{S_5} = I_{S_6} \cdot P_{fail_3}$ is then provided as

$$P_{fail_3} = 1 - \sum_{S_i \in S_N} I_{S_i} = 1 - 2 \left(\frac{\bar{X}_1^2 \exp\left(-\frac{P_{thres}}{\bar{X}_1}\right)}{(\bar{X}_1 + \bar{X}_2)(\bar{X}_1 + \bar{X}_3)} + \frac{\bar{X}_2^2 \exp\left(-\frac{P_{thres}}{\bar{X}_2}\right)}{(\bar{X}_2 + \bar{X}_1)(\bar{X}_2 + \bar{X}_3)} + \frac{\bar{X}_3^2 \exp\left(-\frac{P_{thres}}{\bar{X}_3}\right)}{(\bar{X}_3 + \bar{X}_1)(\bar{X}_3 + \bar{X}_2)} \right). \quad (31)$$

Iterating the procedure, by induction it is proved that P_{fail_n} , the recovery failure probability in the presence of n superimposed signals, is given in closed form by:

$$P_{fail_n} = 1 - \frac{n!}{n} \cdot \sum_{k=1}^n \frac{\bar{X}_k^{n-1} e^{-\frac{P_{thres}}{\bar{X}_k}}}{\prod_{\substack{i=1 \\ i \neq k}}^n (\bar{X}_k + \bar{X}_i)} = 1 - (n-1)! \sum_{k=1}^n \frac{e^{-\frac{P_{thres}}{\bar{X}_k}}}{\prod_{\substack{i=1 \\ i \neq k}}^n \left(1 + \frac{\bar{X}_i}{\bar{X}_k}\right)}. \quad (32)$$

Last expression allows to determine the recovery failure probability in the presence of an arbitrary number of signals in a very effective and quick manner. To this regard, from (11) it is observed that

$$\bar{X}_i = P_{(i)} \bar{\Gamma}_i = \frac{P_{(i)}}{N_0 B} \cdot \bar{h}_i^2 = SNR_{(i)} \cdot \bar{h}_i^2 \quad (33)$$

where $SNR_{(i)}$ is the signal-to-noise ratio of $UE_{(i)}$ and $\bar{h}_i^2$ is the mean of the channel gain h_i^2 . Hence, when the $SNR_{(i)}$ and the $\bar{h}_i^2$ values are provided, the recovery failure probability of uplink power-based NOMA is known. We next assume that the path loss is

$$\bar{h}_i^2 = k_p \cdot D_i^{-\alpha}, \quad (34)$$

where D_i is the distance between the i -th UE and the enodeB, α represents the decay factor and k_p is $k_p = \left(\frac{c}{4\pi f_c}\right)^2$, c being the speed of light, f_c the op-

erating frequency and isotropic antennas are being employed. In this circumstance, (32) specializes to

$$P_{fail_n} = 1 - (n-1)! \sum_{k=1}^n \frac{\exp\left(-\frac{P_{thres}}{SNR_{(k)} k_p D_k^{-\alpha}}\right)}{\prod_{\substack{i=1 \\ i \neq k}}^n \left(1 + \frac{SNR_{(i)}}{SNR_{(k)}} \cdot \left(\frac{D_i}{D_k}\right)^{-\alpha}\right)}. \quad (35)$$

Given P_{thres} is fixed, as well as the operating frequency f_c , the set of distances $D_1, D_2, \dots, D_n$ and the SNR values $SNR_{(1)}, SNR_{(2)}, \dots, SNR_{(n)}$, from (35) the probability of not being able to take advantage of successive interference cancellation is determined right away. Next Section relies on (35) to offer some meaningful insights on the performance of uplink power-based NOMA employed in conjunction with SIC.

IV. NUMERICAL RESULTS

Fig.1 reports P_{fail_2} , the recovery failure probability in the presence of two superimposed signals, as a function of D_1 , the distance of UE_1 from enodeB given in meters, when the second UE , UE_2 , is at the cell edge and the cell radius is $R = 100$ m; $SNR_{(1)}$ takes on different values, namely, $= 15, 12$ and 10 dB, whereas $SNR_{(2)} = 10$ dB, that is to say, a difference of 5, 2 and 0 dB between the transmitted powers of the two UE s is considered. This is in line with the choices performed in [3], where the transmitted powers of two simultaneous users differ for either 5 or 3 dB. The propagation factor is $\alpha = 4$ and two values of the carrier frequency are considered: $f_c = 2$ GHz (solid lines) and 28 GHz (dashed lines). P_{thres} , the minimum difference in received powers is equal to -75 dBm [9]. The frequency effect on the recovery failure probability is evident, highlighting that power-based NOMA is by far more attractive at lower frequencies. Nevertheless, interesting recovery failure probability values can be attained when the distance of UE_1 from its enodeB is small and the gap between $SNR_{(1)}$ and $SNR_{(2)}$ increases. Fig.2 extends the reasoning to the case of three superimposed signals and shows the behavior of the recovery failure probability P_{fail_3} as a function of D_1 for three distinct choices of the $(SNR_{(1)}, SNR_{(2)}, SNR_{(3)})$ triplet, namely: $(10, 10, 10)$ (solid lines), $(12, 10, 10)$ (dashed lines) and $(15, 10, 8)$ (dotted lines), when the carrier frequency is $f_c = 28$ GHz. Different locations of UE_2

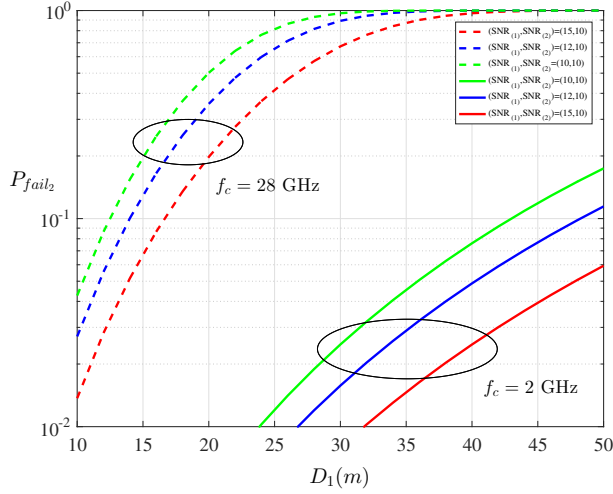


Fig. 1. P_{fail2} as a function of D_1 , $f_c = 2$ and $f_c = 28$ GHz

and UE_3 are examined: $D_2 = 0.2R$ paired with $D_3 = 0.5R$ (red lines), $D_2 = 0.2R$ with $D_3 = 0.7R$ (blue lines), and $D_2 = 0.2R$ with $D_3 = 0.9R$ (green lines). The curves indicate that the influence of D_3 on P_{fail3} is modest: in the examined setting, it is UE_2 that plays the lion's share, being close to enodeB, as $D_2 = 0.2R$ reveals. All curves exhibit a similar shape; however, they become wider for more pronounced differences in the $SNRs$, whose final effect is to increase the range of D_1 values for which the recovery failure probability stays below a predefined threshold (e.g., 10^{-1}). So, the advantage of markedly separating the UEs in terms of SNR , assigning the users with the most favorable channel a higher SNR value, is manifest and numerically quantified. Moreover, note that in the majority of the cases P_{fail3} values fall below $2 \cdot 10^{-1}$.

V. CONCLUSIONS

This paper has identified a novel, analytical method to determine the probability of not being able to take advantage of power-based NOMA on the uplink of a 5G cell, when successive interference cancellation is employed and an arbitrary number of superimposed signals are considered. As a representative example, Rayleigh fading has been examined and the corresponding recovery failure probability provided in closed-form. The dependency of the recovery failure probability on carrier frequency, signal-to-noise ratio of the UEs that are simultaneously transmitting, as well as on their distance from

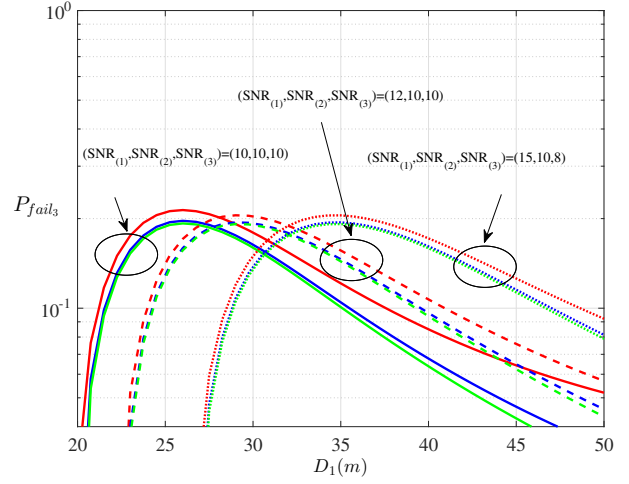


Fig. 2. P_{fail3} as a function of D_1 , $f_c = 28$ GHz

the enodeB has been clearly identified, revealing that even at very high frequencies there exist several operating regions where power-based NOMA combined with SIC exhibits notably low recovery failure probability values, in the presence of two and also three simultaneous users.

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